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TEMPORAL EXPONENTIAL-FAMILY RANDOM GRAPH MODELING (TERGMS) WITH STATNET

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Source for all things STERGM

 Pavel N. Krivitsky and Mark S. Handcock (2014). <u>A Separable</u> <u>Model for Dynamic Networks</u>. *Journal of the Royal Statistical Society, Series B*, Volume 76, Issue 1, pages 29–46.

Terminology

- The phrase "temporal ERGMs," or TERGMs, refers to all ERGMs that are dynamic
- The specific class of TERGMs that have been implemented thus far are called "separable temporal ERGMs," or STERGMs
- In the relevant R package, we left open the possibility that we would develop more in the future

Th	us:

	Cross-sectional	Dynamic
Name of package	ergm	tergm
Name of function in package	ergm	stergm

ERGMs: Review

Probability of observing a graph (set of relationships) y on a fixed set of nodes:

$$P(Y = y | \boldsymbol{\theta}) = \frac{\exp(\boldsymbol{\theta}' \boldsymbol{g}(\boldsymbol{y}))}{k(\boldsymbol{\theta})}$$

Conditional log-odds of a tie

$$logit(P(Y_{ij} = 1 | rest of the graph)) = log\left(\frac{P(Y_{ij} = 1 | rest of the graph)}{P(Y_{ij} = 0 | rest of the graph)}\right)$$
$$= \theta' \partial(g(y))$$

where: g(y) = vector of network statistics $\theta = vector of model parameters$ $k(\theta) = numerator summed over all possible networks on node set y$ $\partial(g(y))$ represents the change in g(y) when Y_{ij} is toggled between 0 and 1



- ERGMs are great for modeling cross-sectional network structure
- But they can only predict the *presence* of a tie; they are unable to separate the processes of *tie formation* and *dissolution*
- Why separate formation from dissolution?

- Intuition: The social forces that facilitate formation of ties are often different from those that facilitate their dissolution.
- Interpretation: Because of this, we would want model parameters to be interpreted in terms of ties formed and ties dissolved.
- Simulation: We want to be able to control cross-sectional network structure and relational durations separately in our disease simulations, matching both to data



- E.g. if a particular type of tie is rare in the cross-section, is that because:
 - They form infrequently?
 - They form frequently, but then dissolve frequently as well?
- The classic approximation formula from epidemiology helps us see the basic relationship among our concepts:



- Core idea:
 - The y_{ij} values (ties in the network) and Y (the set of all y_{ij} values) are now indexed by time
 - Represent evolution from Y_t to Y_{t+1} as a product of two phases: one in which ties are formed and another in which they are dissolved, with each phase a draw from an ERGM.
 - Thus, two formulas: a formation formula and a dissolution formula
 - And, two corresponding sets of statistics

ERGM: Conditional log-odds of a tie existing

 $logit(P(Y_{ij} = 1 | rest of the graph)) = \theta' \partial(g(y))$

STERGM: Conditional log-odds of a tie *forming* (formation model):

 $logit \left(P(Y_{ij,t+1} = 1 | Y_{ij,t} = 0, rest of the graph) \right) = \theta^{+\prime} \partial (g^{+}(y))$

STERGM: Conditional log-odds of a tie *persisting* (dissolution model):

$$logit \left(P(Y_{ij,t+1} = 1 | Y_{ij,t} = 1, rest of the graph) \right) = \theta^{-\prime} \partial (g^{-}(y))$$

where: $g^+(y)$ = vector of network statistics in the formation model θ^+ = vector of parameters in the formation model $g^-(y)$ = vector of network statistics in the dissolution model θ^- = vector of parameters in the dissolution model

Dissolution? Or persistence?

 $logit \left(P(Y_{ij,t+1} = 1 | Y_{ij,t} = 1, rest of the graph) \right) = \theta^{-\prime} \partial (g^{-}(y))$

- The model is expressed as log odds of tie equaling 1 given it equaled 1 at the last time step
- This is done to make it consistent with the formation model, so all the math works out nicely
- But it implies that the model, and thus the coefficients, should be interpreted in terms of effects on relational persistence
- That said, people tend to thing in terms of relational formation and dissolution, since relational dissolution is a more salient event than relational persistence
- Thus, we often use the language of dissolution

During simulation, two processes occur separately within a time step:



- Y⁺ = network in the formation process after evolution
- Y^2 = network in the dissolution process after evolution
- This is the origin of the "S" in STERGM



- The statistical theory in Krivitsky and Handcock 2014:
 - demonstrates a given combination of formation and dissolution model will converge to a stable equilibrium, i.e.:

Prevalence ≈ Incidence x Duration

 This and other work in press provide the statistical theory for methods for estimating the two models, given certain kinds of data

Term = ~edges

	$oldsymbol{ heta}$ /	heta >
Formation model	more new ties created each time step	fewer new ties created each time step
Dissolution (persistence) model	more existing ties pre- served (fewer dissolved); longer average duration	fewer existing ties pre- served (more dissolved); shorter average duration

What combo do you think is most common in empirical networks?

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What combo do you think is most common in empirical social networks?

Term = ~concurrent (# of nodes with degree 2+)

	$\theta \nearrow$	$oldsymbol{ heta}$ >
Formation model	more ties added to actors with exactly 1 tie	fewer ties added to actors with 1 tie
Dissolution (persistence) model	actors with 2 ties more likely to have them be preserved	actors with 2 ties more likely to have them dissolve

What combo do you think is most common in empirical sexual networks?

Term = ~concurrent (# of nodes with degree 2+)

	$\theta \nearrow$	$oldsymbol{ heta}$ >
Formation model	more ties added to actors with exactly 1 tie	fewer ties added to actors with 1 tie
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What combo do you think is most common in empirical sexual networks?

STERGMs: Data sources

1. Multiple cross-sections of complete network data

- easy to work with
- but rare-to-non-existent in some fields
- 2. One snapshot of a cross-sectional network (census, egocentric, or otherwise), plus information on relational durations
 - more common
 - but introduces some statistical issues in estimating relation lengths

STERGMs: nodal dynamics

- All of the statistical theory presented so far regards networks with
 - Dynamic relationships, but still
 - Static actors
- I.e. no births and deaths, no changing of nodal attributes
- The statistical theory of STERGM can handle nodal dynamics during simulation, with a few added tweaks
 - Most important is an offset term to deal with changing population size
 - Without it, density is preserved as population size changes
 - With it, mean degree is preserved as population size changes

STERGMs: nodal dynamics

• For more info, see:

Pavel N. Krivitsky, Mark S. Handcock, and Martina Morris (January 2011). <u>Adjusting for Network Size and Composition Effects in Exponential-Family</u> <u>Random Graph Models</u>. *Statistical Methodology*, 8(4): 319–339

- And for more help with using STERGMs to simulate dynamic networks along with changing nodes and attributes:
 - Take our intensive summer workshop on network modeling for epidemic diffusion
 - Explore the online materials for the workshop (on the statnet webpage)
 - Try the EpiModel package

To the tutorial.....

(reference slides follow)

- In some domains, often takes the form of
 - asking respondents about individual relationships (either with or without identifiers).
 - Often this is the *n* most recent, or all over some time period, or some combination (e.g. up to 3 in the last year)
 - asking whether the relationship is currently ongoing
 - if it's ongoing: asking how long it has been going on (or when it started)
 - if it's over: asking how long it lasted (or when it started and when it ended)
- From this we want to estimate
 - the mean duration of relationships
 - perhaps additional information about the variation in those durations (overall, across categories of respondents, etc.)

Issues?



1. Ongoing durations are right-censored

• can use Kaplan-Meyer or other techniques to deal with

Issues?

2. Relationships are subject to length bias in their probability of being observed

- This can also be adjusted for statistically
- However, complex hybrid inclusion rules (e.g. most recent 3, as long as ongoing at some point in the last year) can make this complicated

- In practice (and for examples in this course), we sometimes rely on an elegant approximation:
 - If relation lengths are approximately exponential/geometric (a big if!), then the effects of length bias and right-censoring cancel out
 - The mean amount of time that the ongoing relationships have lasted until the day of interview (relationship age) is an unbiased estimator of the mean duration of relationships
 - Why?!?

- Exponential/geometric durations suggests a memoryless processes one in which the future does not depend on the past
- Imagine a fair, 6-sided die:

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- 1/6 What is the probability I will get a 1 on my next toss?
- What is the probability I will get a 1 on my next toss given that my previous 1 was five tosses ago?
 - On average, how many tosses will I need before I get my first 1?
 - On average, how many more tosses will I need before I get my next 1, given that my previous 1 was 8 tosses ago?

Geometric	
Parameters	$0 probability (real)$
Support	$k \in \{1, 2, 3, \ldots\}$
Probability mass function (pmf)	$(1-p)^{k-1}p$
Cumulative distribution function (CDF)	$1 - (1 - p)^k$
Mean	1
	$\frac{-}{p}$ 25

- Now, let's imagine this fairly bizarre scenario:
 - You arrive in a room where there are 100 people who have each been flipping one die; they pause when you arrive.
 - You don't know how many sides those dice have, but you know they all have the same number.
 - You are not allowed to ask any information about what they've flipped in the past.
 - The only information people will give you is: how many flips after your arrival does it take until they get their first 1?
 - You are allowed to stay until all of the 100 people get their first 1, and they can inform you of the result.
- Given the information provided you, how will you estimate the number of sizes on the die?

- Simple: when everyone tells you how many flips it takes from your arrival until their first 1, just take the mean of those numbers. Call it m.
- Your best guess for the probability of getting a 1 per flip is 1/m.
- And your best guess for the number of sides is the reciprocal of the probability of any one outcome per flip, which is 1/1/m, which just equals *m* again.
- Voila!

Retrospective relationship surveys are like this, but in reverse:



- If you have something approximating a memoryless process for relational duration, then an unbiased estimator for relationship length is to:
 - ask people about how long their ongoing relationships have lasted up until the present
 - take the mean of that number across respondents.

- In practice, we find that the geometric distribution doesn't often capture the distribution of relational durations overall.
- But, if you divide the relationships into 2+ types, it can do a reasonable job within type
- Especially if you remove any 1-time contacts and model them separately (for populations where they are common)
- Remember: DCMs model pretty much everything as a memoryless process, so approximating one aspect of our model that way is well within common practice